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Evidence from matched employer-employee data and tax-records**

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# The distribution of the gender pay gap in Austria: Evidence from matched employer-employee data and tax records\*

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## Abstract

We examine the gender wage gap in Austria using new matched employer-employee data from 2007. We investigate the gap at the conditional wage distribution of men and women, and decompose it into the parts which are attributed to different characteristics and different returns to these characteristics. We find that women earn on average about 14% less than men for given characteristics, and that about 50% of the gender wage gap cannot be attributed to observable characteristics. The extent of different returns for women and men increase over the wage distribution where wage bargaining is predominantly on an individual basis (in contrast to low wage jobs, where collective bargaining contracts are binding).

*Keywords:* gender wage differentials, quantile regressions, decomposition, matched employer-employee data

*JEL classification:* J31, J71

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# 1 Introduction

Compared to other countries, the gender pay gap in Austria is large. In 2008, it was equal to 25.5% as measured by the EU gender pay gap indicator (European Commission, 2010). The EU gender pay gap indicator measures the average difference between men's and women's gross hourly earnings.<sup>1</sup> Of all EU countries, only the Czech Republic fared worse. To fight this gap, Austrian policy makers required firms to disclose the wages of their workers (Bundeskanzleramt, 2010). Currently, only large firms are covered by the new law, but over time also smaller firms will need to disclose their workers' wages. In addition, affirmative action with countries such as Norway and Spain as role models is often debated. Norway, for example, requires a minimum of 40% of each gender in publicly appointed boards and in all boards of directors in private, shareholder-owned businesses. Spain imposed a similar rule for public sector committees and boards (Bagues and Esteve-Volart, 2007).

The EU gender pay gap indicator, however, does not account for differences between men and women in education, labor market experience or other productivity related variables. To account for these differences, we use a newly constructed data set and decompose the mean wages of men and women in the private and public sector using the technique developed by Blinder (1973) and Oaxaca (1973).<sup>2</sup> This new data set is constructed from administrative files, for the first time matching census information with tax records and social security data, to gauge the reliability of gender wage gap estimates. These new data allow us to go beyond the information typically available in survey data. In particular, we investigate how important it is to account for actual experience, by contrasting the estimated gender pay gaps using exact personal career information with estimates based on the limited information typically available in survey data. In addition, we identify career interruptions and investigate how much estimates of the gender pay gap differ if we ignore the type of interruptions. Another

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<sup>1</sup>For Austria, the gender pay gap is calculated with earnings data from the Structure of Earnings Survey (SES). In contrast to household surveys such as the European Union Survey on Income and Living Conditions (EU-SILC), the SES samples only employees in enterprises with at least ten employees in the private sector (Geisberger and Till, 2009).

<sup>2</sup>There are numerous papers investigating the gender pay gap. For a meta-analysis of studies on the gender pay gap see Weichselbaumer and Winter-Ebmer (2005); for surveys on gender discrimination see Altonji and Black (1999) and Bertrand (2010). In an accompanying paper, we examine the change of the gender wage gap in Austria in the private sector between 2002 and 2007 (Böheim et al., 2011).

set of variables that may explain a part of the gender pay gap are firm-specific variables, which are typically not available in survey data. From our matched employer-employee data, we calculate *inter alia* the ratio of female to male workers within the firm or the ratio for the female to the male median wage in the firm.

The public discussion generally focuses on the average gender pay gap. However, this discussion might be misleading if the gender pay gap differs over the wage distribution. A comprehensive study by Arulampalam et al. (2007) uses the European Community Household Panel, which includes harmonized data on wages and other individual characteristics from various European countries for the years 1995 to 2001. The authors find that in nearly all of the eleven analyzed countries the estimated wage gaps are larger at the top of the distribution than at the bottom of the distribution.<sup>3</sup> We also estimate quantile regressions and contrast the evidence from our estimates with those we obtain from limited information. Here, we follow Melly (2006) and estimate counterfactual distributions, allowing the decomposition of changes in the wage distribution into changes in the regression coefficients, changes in the distribution of covariates and changes in the residuals.

We find that women earn on average about 0.14 log points less than men for given characteristics, and that about 50% of the gender wage gap cannot be attributed to observable characteristics. Differences in returns to characteristics between women and men increase over the wage distribution. This could be attributed to wage bargaining which is predominantly on an individual basis in the high wage segment of the labor market, in contrast to low wage jobs where collective bargaining contracts are the norm. Our findings also demonstrate a remarkable resilience of the estimates. To be sure, the estimates become smaller, the more precise data one has at hands, however, what matters most are good wage information and detailed career information. In our data, the results do hardly change if we account for e.g., the exact number of days on maternity leave—for all practical purposes, an indicator of having been on maternity leave suffices to obtain a reliable estimate of the gender wage gap.

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<sup>3</sup>Other studies have also found significant differences in the gender gap at different quantiles of the log wage distribution. Examples are Albrecht et al. (2003) for Sweden, Fitzenberger and Wunderlich (2001) for the UK, Bonjour and Gerfin (2001) for Switzerland, Gupta et al. (2006) for Denmark, De la Rica et al. (2008) for Spain, Fitzenberger and Wunderlich (2002) for Germany and Albrecht et al. (2009) for the Netherlands.

Firm-specific information additionally contributes to the explanation of the gender pay gap. As can be expected, we find evidence that the firm-specific variables are important wage components and, for example, a higher ratio of female to male workers implies lower wages. We also find that firm-specific variables explain a part of the gender pay gap, for example, the more women work in a firm, the lower is the gender pay gap. If we extend the set of explanatory variables accounting for human capital (education, experience, etc), occupation and industry to include the firm-specific variables as well, the unexplained gender pay gap decreases (in absolute values) from about 0.18 log points to about 0.14 log points, a difference of about 0.04 log points.

## 2 Data and summary statistics

We combine data from Austrian administrative records to construct a new data set to overcome potential weaknesses in earlier studies. Data are from the Austrian General Income Report for 2007, which itself combines data from tax records and the Austrian micro-censuses of 2007, and from the Austrian social security records.<sup>4</sup> An anonymous personal identifier allows the combination of these data, which provides us with data for the analysis of gender wage differences.<sup>5</sup> The merged data contain human capital variables, such as education and experience, workplace characteristics, such as the number of women or the fraction of young workers in a particular workplace, and also complete work histories since 1972, and the sample size corresponds to the number of observations in the micro-censuses.

The Austrian General Income Report, published every other year, provides statistics on the income of all employees, self-employed persons and pensioners in Austria. The Report uses data from tax records; wage data are based on approximately 8.4 million pay slips collected by the Austrian tax authorities and provide information on gross yearly income, paid taxes, paid social contributions and extra compensations. The tax data do not contain information

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<sup>4</sup>The Austrian General Income Report is described in Statistik Austria (2009) and in Rechnungshof (2008). The social security records are described in Zweimüller et al. (2009).

<sup>5</sup>To ensure the anonymity, the actual merging of the data has been handled by an authorized third party. No data that would allow identification of individual persons has been made available to us.

on the number of hours worked and, in addition, taxes are individual data and it is not possible to build household information from the official tax records. For the purposes of the Austrian General Income Report, the tax data are combined with data from the Austrian micro-censuses to generate household level information and to obtain information on e.g., hours worked or formal qualifications. It is therefore an excellent source of information on wage income for employees (Statistik Austria, 2008).

The Austrian micro-census is a quarterly panel survey which collects information on private households. It is representative of the Austrian population and contains information on about 80,000 individuals per year. Every quarter a fifth of the sample is renewed. The micro-census provides information on hours worked, education and detailed information on individual and household characteristics, but it does not contain income information. Combining the information from the micro-census and the tax records allows us to compute exact hourly gross and net wages.

The Austria social security data contain information on individual work experience, tenure and characteristics of the workplace, such as industry or region. A firm's identifier permits the construction of workplace characteristics such as the number of women or the fraction of young workers in a particular workplace. The data also include the reasons for and the length of work interruptions such as unemployment spells or the birth of a child.

Our estimating sample consists of workers who were between 16 and 60 years of age and who worked at least for one hour per week. To account for possible seasonal fluctuations, we restrict our sample to workers who worked for at least 270 days in 2007. Part-time work is defined as working less than 35 hours per week. The sample consists of 4,446 women and 8,919 men who worked in private sector. Table 1 provides summary statistics on our estimating sample. The difference in mean wages in 2007 was almost €4 per hour in the private sector, and women earned on average some 25% less than men, not accounting for differences in characteristics. This "raw" gap was slightly higher than in 1997, when it was about 23.3% (Böheim et al., 2007). Comparing the private sector with the public sector, we see that the difference in mean wages between men and women was smaller in the public than in the

private sector.

Figure 1 plots the distribution of the log hourly wages of men and women in the private sector and Figure 2 includes also full-time employees in the public sector. Both graphs show the same pattern, namely that women's wage distribution is in shape similar to men's, but to the left of it. Women are also slightly less compressed in their wages than men as the peak in their wage distribution is lower than the corresponding peak in the men's distribution. Such distributions are of course only descriptive and do not indicate that women are (unfairly) discriminated against.

One of the most important determinants of the wage is probably the amount of formal education. Women in our sample have on average more formal education than men, however, there are relatively more women who have only compulsory education than men. While women are on average two years younger than men, their average labor market experience difference is about 3.5 years shorter, owing to motherhood and child care responsibilities. Most studies on the gender wage gap can only account for potential experience as the length of and the reasons for work time interruptions is usually not known. Zweimüller and Winter-Ebmer (1994) and Böheim et al. (2007) have demonstrated that it is necessary to account for differences in actual rather than potential experience to obtain reliable estimates of the wage determinants. The summary statistics also show that fewer women than men are married.

The differences in wages might also be related to differences in the workplaces in which women and men work. The summary statistics support such an hypothesis since, for example, women work in smaller workplaces and firms than men, and more women than men work in the public sector. Whether this is the outcome of a selection process or already due to discrimination against women is beyond the scope of the current analysis. However, it should be noted that there are marked differences in the distribution across sectors, for example, the majority of women (27%) works in the retail sector and the majority of men (39%) in manufacturing. We also see that women are concentrated amongst office workers, while men are typically working as craftspersons. Not only do we observe differences in the occupational hierarchy, there is also clear evidence for differences in within-firm hierarchies as merely about

4% of women, in contrast to some 7% of men, have an executive position.

### 3 Methods

As our main tool of analysis we use decomposition techniques and decompose mean wages as well as the wages across the distribution. To decompose mean wages of women and men, we use the technique developed by Blinder (1973) and Oaxaca (1973) and estimate a wage equation for women ( $W$ ) and men ( $M$ ) separately with ordinary least squares:

$$\ln y_i = \beta_i X_i + \epsilon_i, \quad i = W, M, \quad (1)$$

where  $y_i$  is the hourly wage,  $\beta_i$  are the coefficients to be estimated,  $X_i$  is a vector of characteristics, and  $\epsilon_i$  is an i.i.d. error. The difference in the mean wages can be re-written as:

$$\overline{\ln y_M} - \overline{\ln y_W} = \hat{\beta}_M(\overline{X_M} - \overline{X_W}) + (\hat{\beta}_M - \hat{\beta}_W)\overline{X_M}, \quad (2)$$

where  $\hat{\beta}_M(\overline{X_M} - \overline{X_W})$  is difference of the mean characteristics, evaluated at men's prices and  $(\hat{\beta}_M - \hat{\beta}_W)\overline{X_M}$  is the difference in returns to characteristics. The first part is the explained component of the wage difference, i.e., the part which can be ascribed to differences in productivity and the second part is the unexplained component, i.e., the part which cannot be ascribed to differences in productivity. In this way, the difference in mean logarithmic wages is a weighted sum of differences in characteristics and of differences in prices. Equation (1) corresponds to the “male-based” decomposition which assumes that men are paid their marginal product and women are negatively discriminated against. Another way to decompose wages, a “female-based” view, is to assume that women are paid their marginal product and men are positively discriminated against. These two views are limiting cases of the generalized linear



decomposition (Oaxaca and Ransom, 1995):

$$\overline{\ln y_M} - \overline{\ln y_W} = \hat{\beta}^*(\overline{X_M} - \overline{X_W}) + (\hat{\beta}_M - \hat{\beta}^*)\overline{X_M} + (\hat{\beta}^* - \hat{\beta}_F)\overline{X_F}, \quad (3)$$

where  $\hat{\beta}^*$  is a weighted average of the coefficient vectors  $\hat{\beta}_M$  and  $\hat{\beta}_F$ :

$$\hat{\beta}^* = \Omega\hat{\beta}_M + (I - \Omega)\hat{\beta}_F, \quad (4)$$

where  $\Omega$  is a weighting matrix and  $I$  is an identity matrix. The decomposition equations proposed by Blinder (1973) and Oaxaca (1973) represent special cases of the generalized equation in which  $\Omega$  is a null-matrix or equal to  $I$ . Neumark (1988) and Oaxaca and Ransom (1995) estimate a pooled model to derive the counterfactual coefficient vector  $\hat{\beta}^*$ . We follow Reimers (1983) who assumes  $\hat{\beta}^* = \frac{1}{2}\hat{\beta}_M + \frac{1}{2}\hat{\beta}_F$ .

These approaches focus on the mean of the wage distribution and therefore may provide only a limited picture of the differences in wages between women and men. Several authors have found that the mean wage gaps are not representative of the whole distribution. For example, Arulampalam et al. (2007) use the European Community Household Panel, which includes harmonized data on wages and other individual characteristics from various European countries for the years 1995 to 2001. They find that in nearly all of the eleven analyzed countries there is a glass ceiling, i.e., the estimated wage gaps are larger at the top of the distribution than at the bottom of the distribution.

To not only have a look at the the effects of gender and other observables on the conditional mean of the logarithmic wage distribution, but also on different quantiles, we run quantile regressions.<sup>6</sup> Such a regression model specifies the  $q$ -th conditional quantile of the logarithmic wage distribution as a linear function of characteristics:

$$\ln y_{iq} = \beta_{iq}X_i + \epsilon_{iq}, \quad i = W, M, \quad (5)$$

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<sup>6</sup>For an introduction and an overview to quantile regression see Koenker and Bassett (1978) and Koenker and Hallock (2001).

where  $q \in (0, 1)$  and  $\text{Quant}(\epsilon_{iq}|X_i) = 0$ . For each quantile  $q$ , we estimate one equation for women,  $W$ , and men,  $M$ , each. While ordinary least square regressions have the property that the mean of the dependent variable and the mean of the explanatory variables are on the regression line, which makes the decomposition of the dependent variable straightforward, the estimators for the quantile regression models do not have this property. We therefore use a different procedure to calculate the gender wage gap at the  $q$ -the quantile to differences in returns adjusted for characteristics. We follow Melly (2006) and estimate counterfactual distributions, allowing the decomposition of changes in distribution into three factors: changes in regression coefficients, changes in the distribution of covariates and changes in residuals.<sup>7</sup>

## 4 Estimation results

Tables 2 and 3 present the estimated coefficients of two specifications of the gross hourly wage, for full-time men and women in the private sector and in the private and public sector combined. The two specifications differ in the treatment of past labor force statuses, the first specification uses a less detailed measure of past non-employment spells than the second specification. The results do not differ much between specifications, however, the second specification explains slightly more variation in the dependent variable than the first. Thus, we concentrate on the second specification from now on.

The estimated coefficients indicate, for example, that more formal education is associated with higher wages. For example, men with a high school degree earn 37.7 percent more than men with only compulsory schooling. For women, the respective number is equal to 41.0 percent. Men with a university degree earn 50 percent more than men with only compulsory schooling and women earn 47.5 percent more. The estimated coefficients for lower formal education differ considerably between men and women, while differences in secondary and tertiary education are small. The lower estimates for women with apprenticeship may reflect gender-specific educational choices in Austria. Typically, young women choose schools or apprenticeships with social or commercial specialization and young men choose some form of

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<sup>7</sup>We use the Stata code proposed in that paper for our estimates.

technical education. Interestingly, gender-specific specialization also takes place in tertiary education, but here the differences in the returns are lower.

The estimated coefficients further indicate that more experience is also associated with higher wages, whereas unemployment spells, periods of parental leave or times sick are associated with lower wages. In contrast, time in the military has a strong positive effect for both women and men, although this effect is only significant for the latter. We suspect that individuals obtain additional skills like a driver's licence for trucks during their time in the military finally resulting in higher wages.

In addition, wages clearly differ by the type of workplace and they are typically higher in larger workplaces, in banks and in urban areas. We also find statistical evidence for an association between wages and the gender composition of the workplace. Both men and women are estimated to have a significantly lower wage the more women are employed in the firm. This effect is equal to 18.4 percent for women and 22.3 percent for men. However, the ratio of women's to men's wages in the firm is estimated to have a negative relationship with men's wages only, for women we do not find such an association.

#### **4.1 Decomposition results**

Table 4 presents the results from the decomposition of the gender wage gap for employees in the private sector. Note that the sample only contains workers who were working on average at least 35 hours per week on at least 270 days in 2007. In the private sector, women earn on average 23% less than men, without controlling for differences in characteristics. The mean wage difference can be in part explained by differences in characteristics, for example, about 30% of the mean difference can be attributed to differences in formal education or experience. If we use all available information in our data, i.e., controlling for differences in occupation or firm characteristics, we can ascribe almost 50% of the mean wage difference to differences in observable characteristics.

In particular, we find that firm-specific variables contribute to the explanation of the gender pay gap. For example, the more women work in a firm, the lower is the gender pay

gap. Or, if we extend the set of explanatory variables accounting for human capital (education, experience, etc), occupation and industry to include the firm-specific variables as well, the share of the unexplained gender pay gap decreases from about 68% to about 51%. This is a difference of about 17 percentage points. The remainder, the unexplained wage gap, must be ascribed to differential returns to characteristics. Table 5 tabulates the results from the same analysis, using employees from both the private and public sector.

## 4.2 Wage differences over the distribution of wages

The estimated quantile regressions are tabulated in Tables 6 and 7 for the full-time female and male employees in the private sector and in Tables 8 and 9 for female and male full-time employees in both the private and public sectors. For ease of comparison, the first columns in these Tables re-produce the OLS regression results. It is perhaps tedious to compare the estimated coefficients across specifications and a graphical representation, Figure 3, permits a more immediate way of interpreting the results. The blue line decreases over the quantiles of the wage distribution and indicates that wage gap between women and men increases over the wage distribution, however, it is fairly flat over the majority of the distribution. With respect to the explained and unexplained characteristics, we see that at the bottom of the distribution, observed characteristics explain more of the difference than at the intermediate levels (red line).

The differences in returns (green line) are greater at the top of the distribution than at the bottom, which corresponds to an increase of the unexplained part of the wage gap over the wage distribution. This result can be explained by collective bargaining that imposes minimum pay for employment at the bottom of the distribution. Wages at the top of the wage distributions in the private sector are typically the outcome of personal bargaining and thus more flexible. If women have less bargaining power, demand lower wages or are discriminated against, we expect such a distribution of the wage gap. A similar picture is given in Figure 4, where we combine the observations from the private and public sectors.

## 5 Summary and conclusions

We constructed a new data set from administrative sources and decomposed the wages of men and women in Austria. The new data permit a more adequate analysis of the wage differences between women and men. In particular, we use log hourly wages constructed from tax records, employees' characteristics obtained from micro-censuses, and life-time employment histories from social security records. These data allow us to control for differences in formal education, and also in differences in work experiences along with differences in household, workplace, industry or firm characteristics. In contrast to previous research, we also obtained exact measures of experience and work interruptions. All these contribute significantly to the explanation of the gender wage.

Our descriptive analyzes confirm earlier results, women earn on average less than men, they are on average better formally educated than men, but have on average less workplace experience, probably due to child bearing (Böheim et al., 2007; Grünberger and Zulehner, 2009; Pointner and Stiglbauer, 2010). Taking observed differences between women and men into account, by decomposing the mean wage gap, we find that about 50% of the wage gap is due to "fair" discrimination, i.e., observable differences in characteristics. However, the remaining part of the wage gap between women and men cannot be explained by such characteristics. Part of this difference might be caused by unobserved characteristics, e.g., attitude and commitment, however, it is likely that (some of) this difference is caused by unfair discrimination against women.

In order to shed more light on this question, we analyzed the gender wage gap by the quantiles of the wage distributions. We estimated counterfactual distributions, allowing the decomposition of changes in the wage distribution into changes in the regression coefficients, changes in the distribution of covariates and changes in the residuals. Here we find that the wage gap is narrower at the bottom of the distribution than the top, where it is wider. We interpret this widening of the wage gap over the wage distribution as evidence that women fare worse in individual bargaining than men as most low paying jobs are covered by (industry-wide) collective bargaining agreements. In top-paying jobs, individual bargaining is the norm.

In addition, by comparing the public with the private sector, we also see a much narrower gender wage gap in the public sector, which is still dominated by pay scales, than in the private sector.

We can currently only speculate why women fare worse under individual bargaining than men as we lack appropriate data to investigate this question. However, there are several hypotheses, the probably most prominent is that women are unfairly discriminated against. Another explanation is that women, either because of risk-aversion or cultural reasons, bargain for lower wages than men (Bertrand, 2010; Croson and Gneezy, 2009; Riley-Bowles et al., 2005). Given that the gender pay gap at the top of the wage distribution is wide, it warrants to continue to investigate this question.

## References

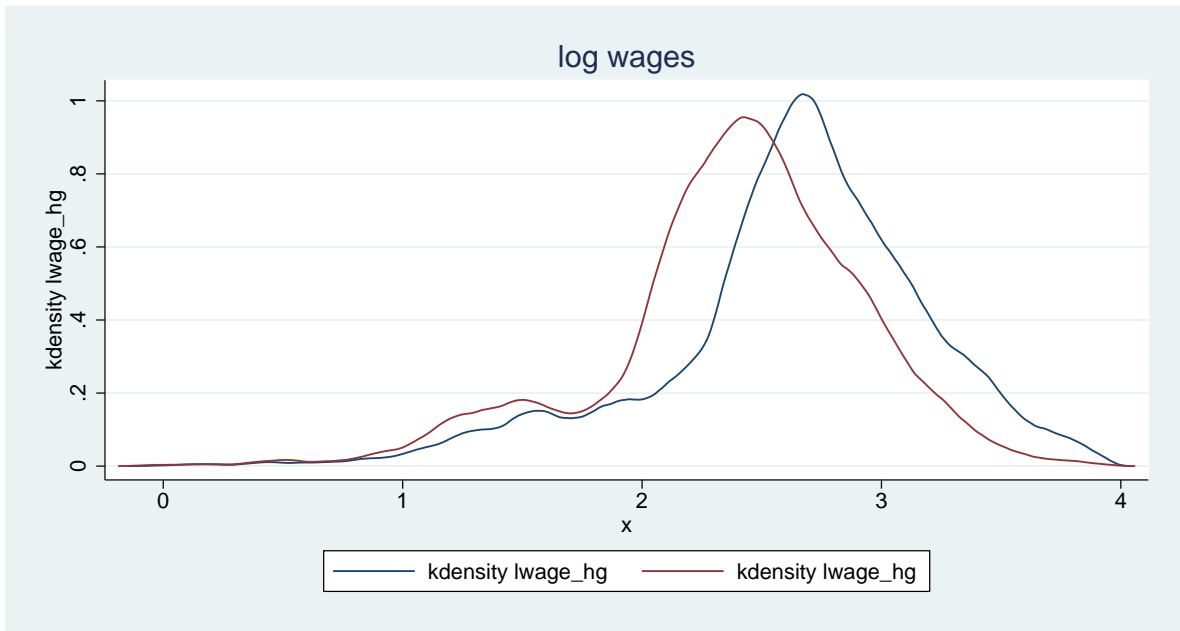
- Albrecht, J., Bjorklund, A., Vroman, S., 2003. Is there a glass ceiling in Sweden? *Journal of Labor Economics* 21 (1), 145–177.
- Albrecht, J., Van Vuuren, A., Vroman, S., 2009. Counterfactual distributions with sample selection adjustments: Econometric theory and an application to the Netherlands. *Labour Economics* 16 (4), 383–396.
- Altonji, J., Black, R., 1999. Race and gender in the labor market. in: O. Ashenfelter and D. Card (eds), *Handbook of Labor Economics*, Vol. 3c, Ch. 48, 3144–3259.
- Arulampalam, W., Booth, A. L., Bryan, M. L., 2007. Is there a glass ceiling over Europe? Exploring the gender pay gap across the wage distribution. *Industrial and Labor Relations Review* 60 (2), 163–186.
- Bagues, M. F., Esteve-Volart, B., 2007. Can gender parity break the glass ceiling? Evidence from a repeated randomized experiment. FEDEA Working Paper 2007-15.
- Bertrand, M., 2010. *Handbook of Labor Economics*. Vol. 4B. North-Holland, Ch. New Perspectives on Gender, pp. 1545–1592.
- Blinder, A. S., 1973. Wage discrimination: Reduced form and structural estimates. *Journal of Human Resources* 18 (4), 436–55.
- Böheim, R., Himpele, K., Mahringer, H., Zulehner, C., 2011. The gender pay gap in Austria: *Tamensi movetur!*, Working Paper.
- Böheim, R., Hofer, H., Zulehner, C., 2007. Wage differences between Austrian men and women: *semper idem?* *Empirica* 34 (3), 213–29.
- Bonjour, D., Gerfin, M., 2001. The unequal distribution of unequal pay—An empirical analysis of the gender wage gap in Switzerland. *Empirical Economics* 26 (2), 407–427.
- Bundeskanzleramt, 2010. Nationaler Aktionsplan Gleichstellung von Frauen und Männern am Arbeitsmarkt. Bundesministerin für Frauen und Öffentlichen Dienst im Bundeskanzleramt.
- Croson, R., Gneezy, U., 2009. Gender Differences in Preferences. *Journal of Economic Literature* 47 (2), 448–474.
- De la Rica, S., Dolado, J., Llorens, V., 2008. Ceilings or floors? Gender wage gaps by education in Spain. *Journal of Population Economics* 21 (3), 751–776.
- European Commission, 2010. Background document on Strategy for Equality between Women and Men 2010-2015.
- Fitzenberger, B., Wunderlich, G., 2001. The Changing Gender Gap Across the Wage Distribution in the U.K. ZEW Discussion Paper, Mannheim 01 (56).
- Fitzenberger, B., Wunderlich, G., 2002. Gender wage differences in West Germany: A cohort analysis. *German Economic Review* 3 (4), 379–414.

- Geisberger, T., Till, M., 2009. Der neue EU-Strukturindikator “gender pay gap”. *Statistische Nachrichten* 1, 64–70.
- Grünberger, K., Zulehner, C., February 2009. Gender-specific wage gaps in Austria. *WIFO Monatsberichte (monthly reports)* 82 (2), 139–50.
- Gupta, N., Oaxaca, R., Smith, N., 2006. Swimming upstream, floating downstream: Comparing women’s relative wage progress in the United States and Denmark. *Industrial and Labor Relations Review* 59 (2), 243–266.
- Koenker, R., Bassett, G., 1978. Regression Quantiles. *Econometrica* 46 (1), 33–50.
- Koenker, R., Hallock, K., 2001. Quantile Regression. *The Journal of Economic Perspectives* 15 (4), 143–156.
- Melly, B., 2006. Estimation of counterfactual distributions using quantile regression, mimeo.
- Neumark, D., 1988. Employers’ discriminatory behavior and the estimation of wage discrimination. *Journal of Human Resources* 22, 279–295.
- Oaxaca, R. L., 1973. Male-female wage differentials in urban labor markets. *International Economic Review* 14, 693–709.
- Oaxaca, R. L., Ransom, M. R., 1995. On discrimination and decomposition of wage differentials. *Journal of Econometrics* 61, 5–21.
- Pointner, W., Stiglbauer, A., May 2010. Changes in the Austrian structure of wages, 1996–2002: Evidence from linked employer-employee data. *Empirica* 37 (2), 105–25.
- Rechnungshof (Ed.), 2008. Bericht des Rechnungshofes gemäß Art. 1 §8 Bezügebegrenzungsgesetz, BGBl. I Nr. 64\1997. 2006 und 2007. Der Rechnungshof, Vienna, Austria.
- Reimers, C. W., 1983. Labor Market Discrimination against Hispanic and Black Men. *Review of Economics and Statistics* 65 (4), 57079.
- Riley-Bowles, H., Babcock, L. C., McGinn, K., 2005. Constraints and Triggers: Situational Mechanics of Gender in Negotiation. *Journal of Personality and Social Psychology* 89 (6), 951–965.
- Statistik Austria (Ed.), 2008. Statistik der Lohnsteuer 2007. Statistik Austria, Vienna.
- Statistik Austria (Ed.), 2009. Standard-Dokumentation. Metainformationen. (Definitionen, Erläuterungen, Methoden, Qualität) zum Allgemeinen Einkommensbericht 2008. Bearbeitungsstand: 25.05.2009. Statistik Austria, Vienna.
- Wechselbaumer, D., Winter-Ebmer, R., 2005. A meta-analysis of the international gender wage gap. *Journal of Economic Surveys* 19 (3), 479–511.
- Zweimüller, J., Winter-Ebmer, R., 1994. Gender wage differentials in private and public sector jobs. *Journal of Population Economics* 7, 271–85.
- Zweimüller, J., Winter-Ebmer, R., Lalive, R., Kuhn, A., Wuellrich, J.-P., Ruf, O., Büchi, S., April 2009. Austrian Social Security Database. Working Paper 0903, NRN: The Austrian Center for Labor Economics and the Analysis of the Welfare State.



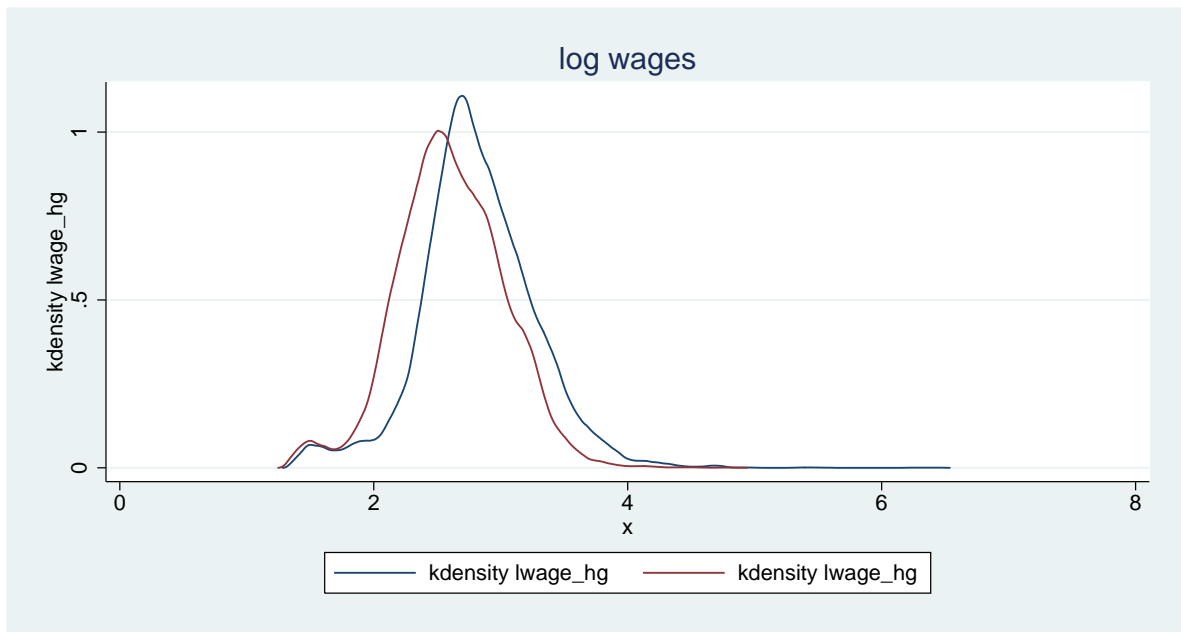
## A Figures

Figure 1: Kernel density of wages in the private sector.



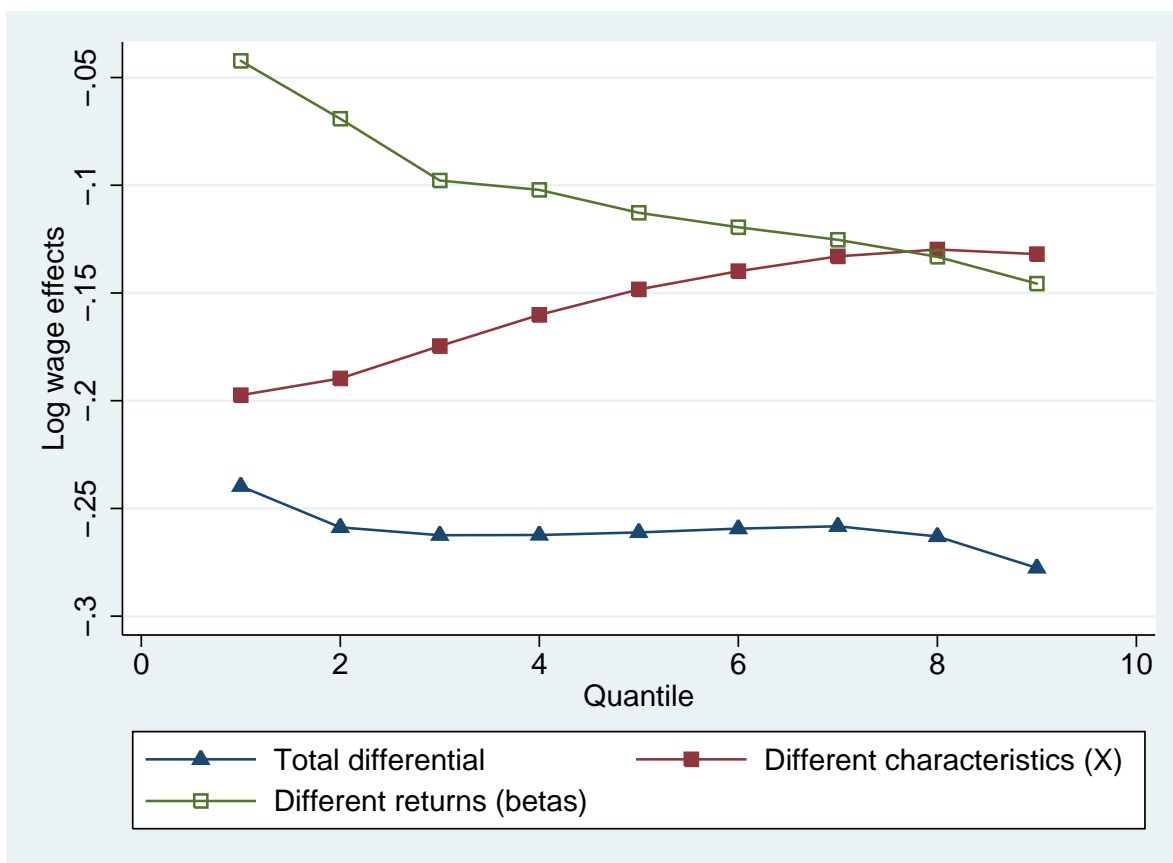
*Note:* 3446 women and 8919 men in private sector employment. Full-time employees only.

Figure 2: Kernel density of wages in the private and public sector.



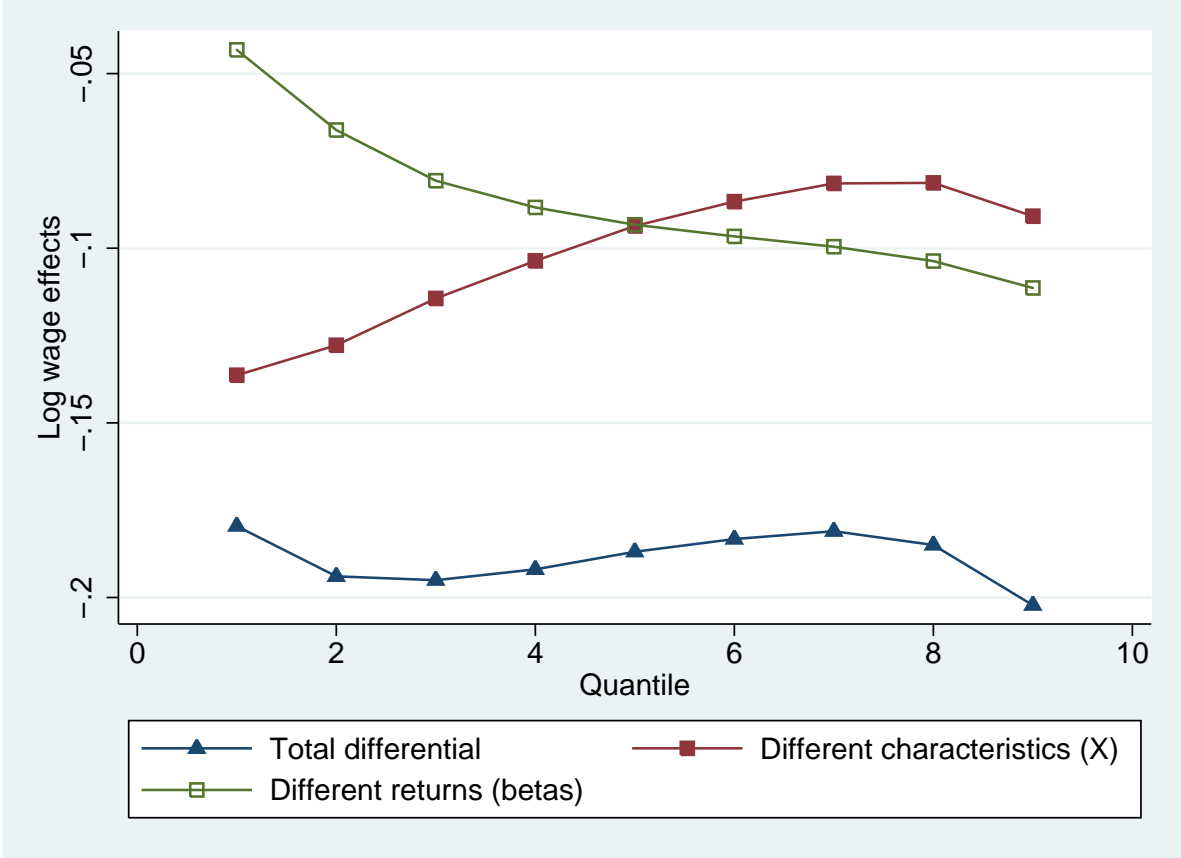
*Note:* 5422 women and 11043 men in private and public sector employment. Full-time employees only.

Figure 3: Quantile decomposition of wages in the private sector.



Note: 3446 women and 8919 men sector employment. Full-time employees only.

Figure 4: Quantile decomposition of wages in the private and public sector.



Note: 5422 women and 11043 men in private and public sector employment. Full-time employees only.

## B Tables

Table 1: Summary statistics, mean (standard deviation).

	Private sector		Private + Public	
	Women	Men	Women	Men
Gross hourly wage (€)	13.120 (5.978)	17.086 (8.018)	14.379 (6.358)	17.503 (7.997)
Education				
Compulsory school	0.230	0.152	0.190	0.137
Apprenticeship	0.344	0.509	0.273	0.482
Secondary school	0.161	0.075	0.201	0.082
High school	0.194	0.136	0.182	0.141
Crafts diploma	0.005	0.065	0.004	0.058
Technical college	0.015	0.012	0.060	0.021
University (Bachelor, Master, MBA)	0.044	0.042	0.077	0.063
University (PhD)	0.007	0.008	0.012	0.015
Age	35.624 (11.625)	37.479 (11.275)	37.646 (11.492)	38.563 (11.169)
Experience	14.823 (9.682)	18.396 (10.467)	15.878 (9.497)	18.893 (10.169)
Tenure (years)	8.030 (7.357)	9.351 (8.582)	9.254 (7.734)	10.274 (8.719)
Length of interruptions (years)	1.004 (1.564)	0.596 (1.035)	0.951 (1.517)	0.557 (1.030)
Length of parental leave	0.382 (0.824)	0.002 (0.051)	0.402 (0.839)	0.003 (0.059)
Time in military	0.000 (0.011)	0.051 (0.176)	0.000 (0.009)	0.047 (0.174)
Time out of labor force	0.140 (0.829)	0.124 (0.635)	0.131 (0.780)	0.125 (0.643)
Time unemployed	0.435 (0.863)	0.349 (0.731)	0.374 (0.814)	0.317 (0.711)
Time sick	0.047 (0.160)	0.070 (0.181)	0.044 (0.158)	0.065 (0.179)
Number of jobs	6.371 (6.641)	7.641 (7.515)	6.273 (6.410)	7.478 (7.513)
Married	0.490	0.650	0.523	0.672
Executive position	0.040	0.073	0.044	0.079
Firm size				
Firm size: 1-10	0.230	0.153	0.204	0.150
Firm size: 11-19	0.146	0.124	0.142	0.123
Firm size: 20-49	0.191	0.188	0.205	0.195
Firm size: 50-499	0.331	0.383	0.338	0.386
Firm size: 500+	0.102	0.153	0.111	0.146
Firm size unknown	0.023	0.027	0.024	0.027
Log workplace size	4.632 (1.891)	4.750 (1.814)	5.295 (2.168)	4.960 (1.930)
Average age of the firm	23.625 (12.594)	24.039 (12.543)	22.907 (12.393)	23.525 (12.281)

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Table 1 – continued from previous page

	Private sector		Private + Public	
	Women	Men	Women	Men
Average age in the firm	37.492 (4.568)	37.679 (4.431)	39.052 (5.038)	38.506 (4.673)
Ratio female to male workers in the firm	0.532 (0.242)	0.266 (0.203)	0.586 (0.224)	0.315 (0.228)
Ratio female to male wage in the firm	0.796 (0.411)	0.762 (0.258)	0.809 (0.354)	0.767 (0.246)
Worker turnover in the firm	41.580 (1050)	4.212 (292)	27.256 (837)	4.792 (264)
Public sector	-	-	0.364	0.192
Occupation				
Soldiers, administrative officers	0.040	0.081	0.034	0.083
Researchers	0.033	0.046	0.119	0.082
Engineers	0.226	0.197	0.253	0.195
Office workers	0.311	0.085	0.251	0.088
Sales	0.184	0.055	0.178	0.083
Craftspersons	0.042	0.308	0.028	0.259
Assembly workers	0.035	0.132	0.026	0.115
Unskilled workers	0.129	0.095	0.111	0.094
Sector				
Agriculture, fishery, mining	0.012	0.014	0.009	0.013
Manufacturing	0.245	0.387	0.157	0.315
Energy, water suppliers, traffic and communication	0.071	0.090	0.048	0.077
Construction	0.040	0.150	0.026	0.125
Whole sale and retail	0.272	0.176	0.174	0.144
Tourism	0.110	0.029	0.071	0.023
Banks, insurance	0.089	0.052	0.056	0.042
Real estate	0.103	0.064	0.067	0.053
Others	0.058	0.036	0.392	0.208
Citizenship				
Austrian	0.917	0.912	0.931	0.923
EU 15	0.018	0.016	0.015	0.015
Others	0.065	0.072	0.054	0.062
Population density				
High	0.344	0.263	0.356	0.276
Medium	0.251	0.269	0.243	0.267
Low	0.405	0.468	0.401	0.457
Region				
Burgenland	0.083	0.083	0.086	0.088
Lower Austria	0.128	0.126	0.130	0.128
Vienna	0.137	0.090	0.147	0.097
Carinthia	0.095	0.100	0.105	0.103
Steiermark	0.100	0.120	0.114	0.120
Upper Austria	0.115	0.148	0.102	0.136
Salzburg	0.121	0.106	0.119	0.111
Tirol	0.108	0.105	0.098	0.104
Vorarlberg	0.113	0.123	0.099	0.115
Number of observations	3,446	8,919	5,422	11,043

Table 2: Estimated wage regressions for the private sector.

	Specification 1		Specification 2	
	Women	Men	Women	Men
Constant	1.615 (0.08)	1.871 (0.05)	1.610 (0.08)	1.832 (0.05)
Education (reference group: compulsory school)				
Apprenticeship	0.189 (0.01)	0.249 (0.01)	0.184 (0.01)	0.217 (0.01)
Secondary school	0.245 (0.02)	0.290 (0.02)	0.242 (0.02)	0.261 (0.02)
High school	0.381 (0.02)	0.442 (0.01)	0.377 (0.02)	0.410 (0.01)
Craftsmen diploma	0.263 (0.07)	0.312 (0.02)	0.263 (0.07)	0.281 (0.02)
Technical college	0.502 (0.04)	0.485 (0.03)	0.500 (0.04)	0.475 (0.03)
University (Bachelor, Master, MBA)	0.604 (0.03)	0.623 (0.02)	0.598 (0.03)	0.603 (0.02)
University (PhD)	0.728 (0.06)	0.711 (0.04)	0.704 (0.06)	0.684 (0.04)
Experience	0.049 (0.00)	0.049 (0.00)	0.051 (0.00)	0.057 (0.00)
Experience squared $\times$ 100	-0.099 (0.01)	-0.098 (0.00)	-0.104 (0.01)	-0.112 (0.00)
Tenure	0.010 (0.00)	0.007 (0.00)	0.009 (0.00)	0.006 (0.00)
Tenure squared $\times$ 100	0.002 (0.01)	0.010 (0.00)	0.005 (0.01)	0.012 (0.00)
Length of interruptions	0.003 (0.01)	0.005 (0.01)		
Length of interruptions $\times$ 100	-0.043 (0.12)	-0.127 (0.10)		
Time unemployed			-0.002 (0.01)	-0.018 (0.01)
Time out of labor force			0.014 (0.01)	0.013 (0.01)
Length of parental leave			-0.009 (0.01)	-0.025 (0.06)
Time in military			0.411 (0.43)	0.259 (0.02)
Time sick			-0.067 (0.03)	-0.041 (0.02)
Married	0.003 (0.01)	0.056 (0.01)	0.006 (0.01)	0.065 (0.01)
Citizenship (reference group: others)				
Austrian citizenship	-0.021 (0.02)	-0.046 (0.01)	-0.019 (0.02)	-0.066 (0.01)
EU15 citizenship	0.052 (0.04)	0.074 (0.03)	0.059 (0.04)	0.091 (0.03)

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Table 2 – continued from previous page

	Specification 1		Specification 2	
	Women	Men	Women	Men
Population density (reference group: high)				
Medium	-0.005	-0.002	-0.006	-0.003
	(0.02)	(0.01)	(0.02)	(0.01)
Low	-0.035	-0.015	-0.037	-0.017
	(0.01)	(0.01)	(0.01)	(0.01)
Region (reference group: Vienna)				
Burgenland	-0.042	0.013	-0.043	0.010
	(0.03)	(0.02)	(0.03)	(0.02)
Lower Austria	-0.007	0.018	-0.010	0.014
	(0.02)	(0.02)	(0.02)	(0.02)
Carinthia	-0.060	-0.026	-0.061	-0.029
	(0.02)	(0.02)	(0.02)	(0.02)
Steiermark	-0.060	-0.013	-0.062	-0.015
	(0.02)	(0.02)	(0.02)	(0.02)
Upper Austria	-0.039	0.036	-0.039	0.031
	(0.02)	(0.02)	(0.02)	(0.02)
Salzburg	-0.010	0.007	-0.014	0.003
	(0.02)	(0.02)	(0.02)	(0.02)
Tirol	-0.008	0.043	-0.010	0.038
	(0.02)	(0.02)	(0.02)	(0.02)
Voralberg	0.017	0.069	0.015	0.061
	(0.02)	(0.02)	(0.02)	(0.01)
Worker status (reference group: white collar)				
Blue collar worker	-0.047	-0.093	-0.045	-0.093
	(0.02)	(0.01)	(0.02)	(0.01)
Occupation (reference group: soldiers, administrative officers)				
Researchers	0.039	-0.015	0.032	-0.014
	(0.04)	(0.02)	(0.04)	(0.02)
Engineers	-0.011	-0.108	-0.013	-0.106
	(0.03)	(0.01)	(0.03)	(0.01)
Office workers	-0.026	-0.136	-0.028	-0.135
	(0.03)	(0.02)	(0.03)	(0.02)
sales	-0.160	-0.254	-0.163	-0.247
	(0.03)	(0.02)	(0.03)	(0.02)
Craftspersons	-0.178	-0.175	-0.181	-0.174
	(0.04)	(0.02)	(0.04)	(0.02)
Assembly workers	-0.120	-0.220	-0.117	-0.219
	(0.04)	(0.02)	(0.04)	(0.02)
Unskilled workers	-0.143	-0.178	-0.142	-0.177
	(0.03)	(0.02)	(0.03)	(0.02)
Industry (reference group: agriculture, fishery, mining)				
Manufacturing	0.128	0.100	0.128	0.100
	(0.05)	(0.03)	(0.05)	(0.03)
Energy, water suppliers, traffic and communication	0.061	0.009	0.060	0.010
	(0.05)	(0.03)	(0.05)	(0.03)
Construction	0.100	0.103	0.099	0.105
	(0.05)	(0.03)	(0.05)	(0.03)
Whole sale and retail	0.078	0.024	0.077	0.024
	(0.04)	(0.03)	(0.04)	(0.03)

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Table 2 – continued from previous page

	Specification 1		Specification 2	
	Women	Men	Women	Men
Tourism	-0.010 (0.05)	-0.044 (0.03)	-0.010 (0.05)	-0.046 (0.03)
Banks, insurance	0.252 (0.05)	0.156 (0.03)	0.251 (0.05)	0.152 (0.03)
Real estate	0.149 (0.05)	0.070 (0.03)	0.147 (0.05)	0.074 (0.03)
Others	0.098 (0.05)	-0.003 (0.03)	0.099 (0.05)	-0.012 (0.03)
Executive position	0.076 (0.03)	0.085 (0.01)	0.077 (0.03)	0.085 (0.01)
Establishment size (reference group: 1-10)				
11-19	0.036 (0.02)	-0.005 (0.01)	0.035 (0.02)	-0.005 (0.01)
20-49	0.065 (0.02)	0.028 (0.01)	0.063 (0.02)	0.029 (0.01)
50-499	0.076 (0.02)	0.054 (0.01)	0.075 (0.02)	0.053 (0.01)
500+	0.160 (0.02)	0.079 (0.01)	0.158 (0.02)	0.078 (0.01)
unknown	-0.009 (0.03)	0.002 (0.02)	-0.008 (0.03)	-0.004 (0.02)
Log Firm size	0.016 (0.00)	0.034 (0.00)	0.016 (0.00)	0.034 (0.00)
Average age in the firm	0.003 (0.00)	0.004 (0.00)	0.003 (0.00)	0.004 (0.00)
Ratio female to male workers in the firm	-0.185 (0.02)	-0.235 (0.02)	-0.184 (0.02)	-0.223 (0.02)
Ratio female to male wages in the firm	0.017 (0.02)	-0.163 (0.02)	0.018 (0.02)	-0.155 (0.02)
Worker turnover in the firm	-0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)
Number of observations	3446	8919	3446	8919
Adjusted R-squared	0.62	0.63	0.62	0.63

*Note:* Ordinary least square regressions. Standard errors in parentheses. 3446 women and 8919 men in private sector employment. Full-time employees only. Specification 1 includes the length of an individual's interruptions in the labor market and its squared, whereas specification 2 replaces these two variables by time unemployed, time out of labor force, length of parental leave, time in military and time sick.

Table 3: Estimated wage regressions for private and public sector.

	Specification 1		Specification 2	
	Women	Men	Women	Men
Constant	1.625 (0.07)	1.848 (0.05)	1.614 (0.07)	1.815 (0.05)
Education (reference group: compulsory school)				
Apprenticeship	0.180 (0.01)	0.230 (0.01)	0.176 (0.01)	0.205 (0.01)
Secondary school	0.256 (0.01)	0.284 (0.01)	0.253 (0.01)	0.261 (0.01)
High school	0.371 (0.01)	0.431 (0.01)	0.368 (0.01)	0.406 (0.01)
Craftsmen diploma	0.281 (0.06)	0.295 (0.01)	0.277 (0.06)	0.272 (0.01)
Technical college	0.415 (0.02)	0.451 (0.02)	0.411 (0.02)	0.434 (0.02)
University (Bachelor, Master, MBA)	0.538 (0.02)	0.612 (0.02)	0.532 (0.02)	0.593 (0.02)
University (PhD)	0.616 (0.04)	0.666 (0.03)	0.603 (0.04)	0.644 (0.03)
Experience	0.045 (0.00)	0.049 (0.00)	0.046 (0.00)	0.056 (0.00)
Experience squared $\times$ 100	-0.086 (0.01)	-0.096 (0.00)	-0.089 (0.01)	-0.109 (0.00)
Tenure	0.008 (0.00)	0.008 (0.00)	0.007 (0.00)	0.006 (0.00)
Tenure squared $\times$ 100	-0.002 (0.01)	0.010 (0.00)	0.000 (0.01)	0.012 (0.00)
Length of interruptions	-0.002 (0.01)	-0.001 (0.01)		
Length of interruptions $\times$ 100	0.028 (0.09)	-0.064 (0.09)		
Time unemployed			-0.000 (0.01)	-0.022 (0.00)
Time out of labor force			0.013 (0.00)	0.013 (0.00)
Length of maternity leave			-0.010 (0.00)	-0.080 (0.05)
Time in military			0.345 (0.42)	0.209 (0.02)
Time sick			-0.062 (0.03)	-0.040 (0.02)
Married	0.006 (0.01)	0.058 (0.01)	0.009 (0.01)	0.065 (0.01)
Citizenship (reference group: others)				
Austria	-0.023 (0.02)	-0.034 (0.01)	-0.022 (0.02)	-0.050 (0.01)
EU15	0.120 (0.04)	0.083 (0.03)	0.125 (0.04)	0.098 (0.03)

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Table 3 – continued from previous page

	Specification 1		Specification 2	
	Women	Men	Women	Men
Population density (reference group: high)				
Medium	-0.012	-0.004	-0.013	-0.005
	(0.01)	(0.01)	(0.01)	(0.01)
Low	-0.023	-0.018	-0.023	-0.020
	(0.01)	(0.01)	(0.01)	(0.01)
Region (reference group: Vienna)				
Burgenland	-0.031	0.013	-0.031	0.010
	(0.02)	(0.02)	(0.02)	(0.02)
Lower Austria	-0.001	0.014	-0.003	0.011
	(0.02)	(0.01)	(0.02)	(0.01)
Carinthia	-0.039	-0.024	-0.040	-0.027
	(0.02)	(0.01)	(0.02)	(0.01)
Steiermark	-0.028	-0.013	-0.029	-0.014
	(0.02)	(0.01)	(0.02)	(0.01)
Upper Austria	-0.023	0.032	-0.023	0.026
	(0.02)	(0.01)	(0.02)	(0.01)
Salzburg	-0.003	0.003	-0.005	-0.002
	(0.02)	(0.01)	(0.02)	(0.01)
Tirol	0.008	0.034	0.006	0.029
	(0.02)	(0.01)	(0.02)	(0.01)
Vorarlberg	0.032	0.065	0.031	0.058
	(0.02)	(0.01)	(0.02)	(0.01)
Worker status (reference group: white collar)				
Blue collar worker	-0.070	-0.092	-0.069	-0.092
	(0.01)	(0.01)	(0.01)	(0.01)
Civil servants	0.015	-0.006	0.014	-0.013
	(0.02)	(0.01)	(0.02)	(0.01)
Other public sector employees	-0.071	-0.115	-0.070	-0.117
	(0.02)	(0.02)	(0.02)	(0.02)
Occupation (reference group: soldiers, administrative officers)				
Researchers	0.038	-0.016	0.037	-0.013
	(0.03)	(0.02)	(0.03)	(0.02)
Engineers	-0.044	-0.080	-0.044	-0.076
	(0.02)	(0.01)	(0.02)	(0.01)
Office workers	-0.088	-0.110	-0.089	-0.107
	(0.02)	(0.01)	(0.02)	(0.01)
Sales	-0.186	-0.183	-0.186	-0.177
	(0.02)	(0.02)	(0.02)	(0.02)
Craftspersons	-0.228	-0.150	-0.229	-0.147
	(0.03)	(0.01)	(0.03)	(0.01)
Assembly workers	-0.158	-0.196	-0.154	-0.192
	(0.03)	(0.02)	(0.03)	(0.02)
Unskilled workers	-0.178	-0.153	-0.175	-0.150
	(0.03)	(0.02)	(0.03)	(0.02)
Industry (reference group: agriculture, fishery, mining)				
Manufacturing	0.123	0.124	0.124	0.124
	(0.04)	(0.03)	(0.04)	(0.03)
Energy, water suppliers, traffic and communication	0.065	0.033	0.065	0.035
	(0.04)	(0.03)	(0.04)	(0.03)

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Table 3 – continued from previous page

	Specification 1		Specification 2	
	Women	Men	Women	Men
Construction	0.103 (0.05)	0.123 (0.03)	0.104 (0.05)	0.125 (0.03)
Whole sale and retail	0.057 (0.04)	0.040 (0.03)	0.057 (0.04)	0.039 (0.03)
Tourism	-0.034 (0.04)	-0.060 (0.03)	-0.033 (0.04)	-0.060 (0.03)
Banks, insurance	0.252 (0.04)	0.173 (0.03)	0.253 (0.04)	0.169 (0.03)
Real estate	0.144 (0.04)	0.086 (0.03)	0.142 (0.04)	0.090 (0.03)
Others	0.095 (0.04)	0.008 (0.03)	0.095 (0.04)	0.005 (0.03)
Executive position	0.117 (0.02)	0.092 (0.01)	0.118 (0.02)	0.093 (0.01)
Establishment size (reference group: 1-10)				
11-19	0.050 (0.01)	0.001 (0.01)	0.050 (0.01)	0.001 (0.01)
20-49	0.065 (0.01)	0.036 (0.01)	0.064 (0.01)	0.037 (0.01)
50-499	0.091 (0.01)	0.059 (0.01)	0.090 (0.01)	0.059 (0.01)
500+	0.153 (0.02)	0.089 (0.01)	0.152 (0.02)	0.089 (0.01)
unknown	-0.028 (0.03)	0.003 (0.02)	-0.025 (0.03)	-0.001 (0.02)
Log Firm size	0.017 (0.00)	0.028 (0.00)	0.017 (0.00)	0.028 (0.00)
Average age in the firm	0.005 (0.00)	0.004 (0.00)	0.005 (0.00)	0.004 (0.00)
Ratio female to male workers in the firm	-0.164 (0.02)	-0.221 (0.02)	-0.162 (0.02)	-0.208 (0.02)
Ratio female to male wages in the firm	0.024 (0.01)	-0.179 (0.02)	0.024 (0.01)	-0.171 (0.02)
Worker turnover in the firm	-0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)
Number of observations	5422	11043	5422	11043
Adjusted R-squared	0.64	0.62	0.64	0.63

*Note:* Ordinary least square regressions. Standard errors in parentheses. 5422 women and 11043 men in private and public sector employment. Full-time employees only. Specification 1 includes the length of an individual's interruptions in the labor market and its squared, whereas specification 2 replaces these two variables by time unemployed, time out of labor force, length of parental leave, time in military and time sick.

Table 4: Decompositions of the wage difference for the private sector.

	(1)	(2)	(3)	(4)
Differences in observed characteristics, $\Delta X$ in % of the raw gap (23.4%)		0.070 30,0	0.075 32,0	0.120 49,0
Differences in returns, $\Delta \hat{\beta}$ (Reimers, 1983) in % of the raw gap (23.4%)	0,255	0,185 70,0	0,180 68,0	0,135 51,0
Education, experience, tenure, interruptions, family status, citizenship, region, density		x	x	x
Worker status, occupation, industry			x	x
Establishment size, firm characteristics, hierarchy				x
Number of observations	12,365	12,365	12,365	12,365

*Note:* Blinder-Oaxaca decompositions. 3446 women and 8919 men in private sector employment. Full-time employees only. For the decomposition three specifications based on specification 1 as depicted in Table 2 are used: In column (2), the independent variables are education, experience, experience squared, tenure, tenure squared, interruptions, interruptions squared, family status, citizenship, dummy variables for regions and population density. In column (3), we add dummy variables for worker status, occupation and industry and in column (4), we add dummy variables for establishment size, logarithm of firm size, average age of workers in the firm, Ratio female to male workers in the firm, ratio female to male wages in the firm, worker turnover in the firm and a dummy variable for a leading position.

Table 5: Decompositions of the wage difference for the private and public sector.

	(1)	(2)	(3)	(4)
Differences in observed characteristics, $\Delta X$ in % of the raw gap (18.1%)		0.070 15,2	0.075 17,8	0.120 38,7
Differences in returns, $\Delta \hat{\beta}$ (Reimers, 1983) in % of the raw gap (18.1%)	0,191	0,162 84,8	0,157 82,2	0,117 61,3
Education, experience, interruptions, family status, citizenship, region, density		x	x	x
Worker status, occupation, industry			x	x
Establishment size, firm characteristics, hierarchy				x
Number of observations	16,465	16,465	16,465	16,465

*Note:* Blinder-Oaxaca decompositions. 5422 women and 11043 men in private and public sector employment. Full-time employees only. For the decomposition three specifications based on specification 1 as depicted in Table 3 are used: In column (2), the independent variables are education, experience, experience squared, tenure, tenure squared, interruptions, interruptions squared, family status, citizenship, dummy variables for regions and population density. In column (3), we add dummy variables for worker status, occupation and industry and in column (4), we add dummy variables for establishment size, logarithm of firm size, average age of workers in the firm, Ratio female to male workers in the firm, ratio female to male wages in the firm, worker turnover in the firm and a dummy variable for a leading position.

Table 6: OLS and Quantile regressions for women in the private sector.

	OLS	10%	25%	50%	75%	90%
Constant	1.615 (0.08)	1.314 (0.15)	1.387 (0.09)	1.800 (0.08)	2.046 (0.09)	2.353 (0.10)
Education (reference group: compulsory school)						
Apprenticeship	0.189 (0.01)	0.227 (0.03)	0.222 (0.02)	0.141 (0.01)	0.115 (0.02)	0.108 (0.02)
Secondary school	0.245 (0.02)	0.261 (0.03)	0.277 (0.02)	0.212 (0.02)	0.190 (0.02)	0.194 (0.02)
High school	0.381 (0.02)	0.441 (0.04)	0.420 (0.02)	0.327 (0.02)	0.311 (0.02)	0.291 (0.02)
Craftsmen diploma	0.263 (0.07)	0.228 (0.12)	0.345 (0.08)	0.192 (0.07)	0.218 (0.08)	0.168 (0.07)
Technical college	0.502 (0.04)	0.493 (0.08)	0.521 (0.05)	0.477 (0.04)	0.397 (0.05)	0.394 (0.05)
University (Bachelor, Master, MBA)	0.604 (0.03)	0.583 (0.05)	0.633 (0.04)	0.538 (0.03)	0.554 (0.04)	0.517 (0.04)
University (PhD)	0.728 (0.06)	0.758 (0.11)	0.795 (0.07)	0.617 (0.06)	0.659 (0.08)	0.626 (0.08)
Experience	0.049 (0.00)	0.058 (0.00)	0.049 (0.00)	0.044 (0.00)	0.040 (0.00)	0.040 (0.00)
Experience squared $\times$ 100	-0.099 (0.01)	-0.122 (0.01)	-0.101 (0.01)	-0.090 (0.01)	-0.076 (0.01)	-0.074 (0.01)
Tenure	0.010 (0.00)	0.009 (0.00)	0.012 (0.00)	0.008 (0.00)	0.007 (0.00)	0.007 (0.00)
Tenure squared $\times$ 100	0.002 (0.01)	0.010 (0.01)	-0.002 (0.01)	0.010 (0.01)	0.012 (0.01)	0.011 (0.01)
Length of interruptions	0.003 (0.01)	0.003 (0.01)	0.001 (0.01)	0.002 (0.01)	0.008 (0.01)	0.010 (0.01)
Length of interruptions $\times$ 100	-0.043 (0.12)	0.006 (0.19)	0.001 (0.12)	-0.047 (0.11)	-0.118 (0.13)	-0.174 (0.15)
Married	0.003 (0.01)	-0.018 (0.02)	0.002 (0.01)	0.008 (0.01)	0.010 (0.01)	0.003 (0.02)
Citizenship (reference group: others)						
Austrian	-0.021 (0.02)	-0.012 (0.04)	0.025 (0.02)	0.017 (0.02)	-0.033 (0.03)	-0.050 (0.03)
EU15	0.052 (0.04)	0.085 (0.07)	0.071 (0.05)	0.029 (0.04)	0.031 (0.05)	0.044 (0.06)
Population density (reference group: high)						
Medium	-0.005 (0.02)	-0.010 (0.03)	-0.036 (0.02)	-0.004 (0.01)	0.003 (0.02)	0.024 (0.02)
Low	-0.035 (0.01)	-0.027 (0.03)	-0.050 (0.02)	-0.041 (0.01)	-0.022 (0.02)	-0.022 (0.02)
Worker status (reference group: white collar)						
Blue collar worker	-0.047 (0.02)	-0.065 (0.04)	-0.076 (0.02)	-0.070 (0.02)	-0.077 (0.02)	-0.083 (0.02)
Executive position	0.076 (0.03)	0.090 (0.05)	0.066 (0.03)	0.071 (0.03)	0.109 (0.03)	0.113 (0.04)
Establishment size (reference group: 1-10)						

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Table 6 – continued from previous page

	OLS	10%	25%	50%	75%	90%
11-19	0.036 (0.02)	0.044 (0.03)	0.043 (0.02)	0.020 (0.02)	0.021 (0.02)	0.023 (0.02)
20-49	0.065 (0.02)	0.071 (0.03)	0.063 (0.02)	0.045 (0.02)	0.026 (0.02)	0.019 (0.02)
50-499	0.076 (0.02)	0.098 (0.03)	0.069 (0.02)	0.065 (0.01)	0.032 (0.02)	0.035 (0.02)
500+	0.160 (0.02)	0.164 (0.04)	0.164 (0.02)	0.149 (0.02)	0.099 (0.03)	0.113 (0.03)
unknown	-0.009 (0.03)	-0.048 (0.06)	-0.009 (0.04)	-0.031 (0.03)	-0.027 (0.04)	-0.041 (0.04)
Log Firm size	0.016 (0.00)	0.017 (0.01)	0.014 (0.00)	0.012 (0.00)	0.019 (0.00)	0.017 (0.00)
Average age in the firm	0.003 (0.00)	0.002 (0.00)	0.003 (0.00)	0.001 (0.00)	0.001 (0.00)	-0.001 (0.00)
Ratio female to male workers in the firm	-0.185 (0.02)	-0.166 (0.04)	-0.172 (0.03)	-0.169 (0.02)	-0.197 (0.03)	-0.233 (0.03)
Ratio female to male wages in the firm	0.017 (0.02)	-0.050 (0.03)	-0.017 (0.02)	0.036 (0.02)	0.084 (0.02)	0.095 (0.02)
Worker turnover in the firm	-0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	0.000 (0.00)
Number of observations	3446	3446	3446	3446	3446	3446

*Note:* Standard errors in parentheses. 3446 women in private sector employment. Full-time employees only. Specification 1 as depicted in Table 2 is used. All regressions include region, occupation and industry effects.



Table 7: OLS and Quantile regressions for men in the private sector.

	OLS	10%	25%	50%	75%	90%
Constant	1.871 (0.05)	1.395 (0.08)	1.628 (0.06)	1.999 (0.06)	2.307 (0.05)	2.375 (0.07)
Education (reference group: compulsory school)						
Apprenticeship	0.249 (0.01)	0.363 (0.01)	0.309 (0.01)	0.212 (0.01)	0.165 (0.01)	0.140 (0.01)
Secondary school	0.290 (0.02)	0.361 (0.02)	0.327 (0.02)	0.269 (0.02)	0.222 (0.02)	0.187 (0.02)
High school	0.442 (0.01)	0.574 (0.02)	0.509 (0.02)	0.393 (0.01)	0.330 (0.02)	0.314 (0.02)
Craftsmen diploma	0.312 (0.02)	0.429 (0.02)	0.363 (0.02)	0.284 (0.02)	0.223 (0.02)	0.196 (0.02)
Technical college	0.485 (0.03)	0.512 (0.05)	0.540 (0.04)	0.484 (0.03)	0.392 (0.04)	0.342 (0.04)
University (Bachelor, Master, MBA)	0.623 (0.02)	0.710 (0.04)	0.669 (0.02)	0.579 (0.02)	0.556 (0.02)	0.550 (0.03)
University (PhD)	0.711 (0.04)	0.717 (0.06)	0.704 (0.04)	0.664 (0.04)	0.698 (0.04)	0.704 (0.05)
Experience	0.049 (0.00)	0.055 (0.00)	0.053 (0.00)	0.045 (0.00)	0.039 (0.00)	0.037 (0.00)
Experience squared $\times$ 100	-0.098 (0.00)	-0.109 (0.01)	-0.107 (0.00)	-0.091 (0.00)	-0.076 (0.00)	-0.069 (0.01)
Tenure	0.007 (0.00)	0.009 (0.00)	0.006 (0.00)	0.003 (0.00)	0.003 (0.00)	0.004 (0.00)
Tenure squared $\times$ 100	0.010 (0.00)	0.007 (0.01)	0.012 (0.01)	0.022 (0.00)	0.020 (0.01)	0.017 (0.01)
Length of interruptions	0.005 (0.01)	0.017 (0.01)	0.003 (0.01)	-0.008 (0.01)	-0.016 (0.01)	-0.018 (0.01)
Length of interruptions $\times$ 100	-0.127 (0.10)	-0.225 (0.12)	-0.086 (0.09)	-0.071 (0.10)	0.007 (0.12)	0.066 (0.14)
Married	0.056 (0.01)	0.054 (0.01)	0.058 (0.01)	0.043 (0.01)	0.040 (0.01)	0.038 (0.01)
Citizenship (reference group: others)						
Austrian	-0.046 (0.01)	-0.026 (0.02)	-0.033 (0.02)	-0.021 (0.01)	-0.016 (0.01)	0.000 (0.02)
EU15	0.074 (0.03)	0.068 (0.05)	0.047 (0.03)	0.037 (0.03)	0.066 (0.03)	0.141 (0.04)
Population density (reference group: high)						
Medium	-0.002 (0.01)	-0.007 (0.02)	0.001 (0.01)	-0.000 (0.01)	-0.013 (0.01)	-0.020 (0.01)
Low	-0.015 (0.01)	-0.029 (0.02)	-0.023 (0.01)	-0.017 (0.01)	-0.023 (0.01)	-0.019 (0.01)
Worker status (reference group: white collar)						
Blue collar worker	-0.093 (0.01)	-0.095 (0.02)	-0.085 (0.01)	-0.099 (0.01)	-0.125 (0.01)	-0.147 (0.01)
Executive position	0.085 (0.01)	0.039 (0.02)	0.070 (0.02)	0.098 (0.01)	0.120 (0.02)	0.123 (0.02)
Establishment size (reference group: 1-10)						
11-19	-0.005 (0.01)	-0.001 (0.02)	-0.002 (0.02)	0.005 (0.01)	-0.024 (0.02)	-0.037 (0.02)

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Table 7 – continued from previous page

	OLS	10%	25%	50%	75%	90%
20-49	(0.01) 0.028	(0.02) 0.024	(0.01) 0.028	(0.01) 0.029	(0.01) 0.013	(0.02) 0.005
50-499	(0.01) 0.054	(0.02) 0.036	(0.01) 0.040	(0.01) 0.056	(0.01) 0.046	(0.02) 0.037
500+	(0.01) 0.079	(0.02) 0.066	(0.01) 0.083	(0.01) 0.083	(0.01) 0.065	(0.01) 0.041
unknown	(0.01) 0.002	(0.02) 0.031	(0.02) 0.018	(0.02) -0.009	(0.02) -0.007	(0.02) 0.026
Log Firm size	(0.02) 0.034	(0.03) 0.039	(0.02) 0.036	(0.02) 0.031	(0.02) 0.029	(0.03) 0.029
Average age in the firm	(0.00) 0.004	(0.00) 0.002	(0.00) 0.003	(0.00) 0.004	(0.00) 0.006	(0.00) 0.006
Ratio female to male workers in the firm	(0.00) -0.235	(0.00) -0.251	(0.00) -0.261	(0.00) -0.287	(0.00) -0.240	(0.00) -0.231
Ratio female to male wages in the firm	(0.02) -0.163	(0.03) -0.175	(0.02) -0.149	(0.02) -0.131	(0.02) -0.118	(0.03) -0.119
Worker turnover in the firm	(0.02) -0.000	(0.03) 0.000	(0.02) -0.000	(0.02) -0.000	(0.02) -0.000	(0.02) -0.000
Number of observations	(0.00) 8919	(0.00) 8919	(0.00) 8919	(0.00) 8919	(0.00) 8919	(0.00) 8919

*Note:* Standard errors in parentheses. 8919 men in private sector employment. Full-time employees only. Specification 1 as depicted in Table 2 is used. All regressions include region, occupation and industry effects.

Table 8: OLS and Quantile regressions for women in the private and public sector.

	OLS	10%	25%	50%	75%	90%
Constant	1.625 (0.07)	1.088 (0.13)	1.367 (0.09)	1.740 (0.06)	2.081 (0.07)	2.366 (0.09)
Education (reference group: compulsory school)						
Apprenticeship	0.180 (0.01)	0.229 (0.02)	0.204 (0.02)	0.137 (0.01)	0.099 (0.01)	0.078 (0.02)
Secondary school	0.256 (0.01)	0.290 (0.03)	0.283 (0.02)	0.230 (0.01)	0.192 (0.01)	0.182 (0.02)
High school	0.371 (0.01)	0.433 (0.03)	0.402 (0.02)	0.326 (0.01)	0.291 (0.02)	0.259 (0.02)
Craftsmen diploma	0.281 (0.06)	0.283 (0.09)	0.325 (0.08)	0.244 (0.05)	0.250 (0.07)	0.205 (0.09)
Technical college	0.415 (0.02)	0.477 (0.04)	0.458 (0.03)	0.397 (0.02)	0.328 (0.03)	0.324 (0.04)
University (Bachelor, Master, MBA)	0.538 (0.02)	0.551 (0.04)	0.544 (0.03)	0.502 (0.02)	0.500 (0.02)	0.475 (0.03)
University (PhD)	0.616 (0.04)	0.619 (0.07)	0.578 (0.05)	0.540 (0.03)	0.562 (0.04)	0.530 (0.06)
Experience	0.045 (0.00)	0.049 (0.00)	0.044 (0.00)	0.039 (0.00)	0.039 (0.00)	0.036 (0.00)
Experience squared $\times$ 100	-0.086 (0.01)	-0.101 (0.01)	-0.086 (0.01)	-0.076 (0.00)	-0.072 (0.01)	-0.065 (0.01)
Tenure	0.008 (0.00)	0.013 (0.00)	0.010 (0.00)	0.008 (0.00)	0.005 (0.00)	0.005 (0.00)
Tenure squared $\times$ 100	-0.002 (0.01)	-0.011 (0.01)	-0.005 (0.01)	0.000 (0.01)	0.011 (0.01)	0.010 (0.01)
Length of interruptions	-0.002 (0.01)	0.007 (0.01)	-0.001 (0.01)	-0.006 (0.01)	-0.002 (0.01)	-0.002 (0.01)
Length of interruptions $\times$ 100	0.028 (0.09)	-0.009 (0.18)	0.028 (0.13)	0.076 (0.08)	-0.015 (0.10)	-0.065 (0.12)
Married	0.006 (0.01)	-0.015 (0.01)	-0.003 (0.01)	0.012 (0.01)	0.009 (0.01)	0.004 (0.01)
Citizenship (reference group: others)						
Austrian	-0.023 (0.02)	-0.001 (0.03)	0.010 (0.02)	0.006 (0.02)	-0.034 (0.02)	-0.026 (0.03)
EU15	0.120 (0.04)	0.136 (0.06)	0.111 (0.05)	0.098 (0.03)	0.116 (0.04)	0.121 (0.05)
Population density (reference group: high)						
Medium	-0.012 (0.01)	-0.033 (0.02)	-0.023 (0.02)	0.002 (0.01)	0.005 (0.01)	0.015 (0.02)
Low	-0.023 (0.01)	-0.025 (0.02)	-0.023 (0.01)	-0.015 (0.01)	-0.012 (0.01)	-0.009 (0.02)
Worker status (reference group: white collar)						
Blue collar worker	-0.070 (0.01)	-0.063 (0.03)	-0.091 (0.02)	-0.080 (0.01)	-0.090 (0.01)	-0.121 (0.02)
Civil servants	0.015 (0.02)	0.062 (0.03)	0.034 (0.02)	0.013 (0.01)	0.009 (0.02)	0.003 (0.03)
Other public sector employees	-0.071 (0.02)	-0.045 (0.03)	-0.066 (0.02)	-0.068 (0.01)	-0.089 (0.02)	-0.093 (0.03)

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Table 8 – continued from previous page

	OLS	10%	25%	50%	75%	90%
Executive position	0.117 (0.02)	0.110 (0.03)	0.104 (0.02)	0.117 (0.01)	0.112 (0.02)	0.130 (0.02)
Establishment size (reference group: 1-10)						
11-19	0.050 (0.01)	0.057 (0.02)	0.048 (0.02)	0.037 (0.01)	0.043 (0.01)	0.042 (0.02)
20-49	0.065 (0.01)	0.069 (0.02)	0.061 (0.02)	0.047 (0.01)	0.057 (0.01)	0.055 (0.02)
50-499	0.091 (0.01)	0.114 (0.02)	0.086 (0.02)	0.078 (0.01)	0.071 (0.01)	0.076 (0.02)
500+	0.153 (0.02)	0.160 (0.03)	0.140 (0.02)	0.129 (0.01)	0.132 (0.02)	0.146 (0.02)
unknown	-0.028 (0.03)	-0.082 (0.04)	-0.033 (0.03)	-0.045 (0.02)	-0.026 (0.03)	-0.031 (0.04)
Log Firm size	0.017 (0.00)	0.023 (0.00)	0.019 (0.00)	0.014 (0.00)	0.015 (0.00)	0.012 (0.00)
Average age in the firm	0.005 (0.00)	0.006 (0.00)	0.004 (0.00)	0.002 (0.00)	0.002 (0.00)	0.001 (0.00)
Ratio female to male workers in the firm	-0.164 (0.02)	-0.116 (0.04)	-0.143 (0.03)	-0.138 (0.02)	-0.170 (0.02)	-0.222 (0.03)
Ratio female to male wages in the firm	0.024 (0.01)	-0.010 (0.03)	0.006 (0.02)	0.052 (0.01)	0.080 (0.01)	0.087 (0.02)
Worker turnover in the firm	-0.000 (0.00)	0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	0.000 (0.00)
Number of observations	5422	5422	5422	5422	5422	5422

*Note:* Standard errors in parentheses. 5422 women in private and public sector employment. Full-time employees only. Specification 1 as depicted in Table 3 is used. All regressions include region, occupation and industry effects.

Table 9: OLS and Quantile regressions for men in the private and public sector.

	OLS	10%	25%	50%	75%	90%
Constant	1.848 (0.05)	1.309 (0.07)	1.569 (0.06)	1.955 (0.06)	2.266 (0.06)	2.360 (0.06)
Education (reference group: compulsory school)						
Apprenticeship	0.230 (0.01)	0.323 (0.01)	0.282 (0.01)	0.194 (0.01)	0.145 (0.01)	0.122 (0.01)
Secondary school	0.284 (0.01)	0.357 (0.02)	0.328 (0.02)	0.260 (0.02)	0.212 (0.02)	0.175 (0.02)
High school	0.431 (0.01)	0.552 (0.02)	0.488 (0.01)	0.384 (0.01)	0.323 (0.01)	0.299 (0.02)
Craftsmen diploma	0.295 (0.01)	0.394 (0.02)	0.341 (0.02)	0.264 (0.02)	0.210 (0.02)	0.174 (0.02)
Technical college	0.451 (0.02)	0.542 (0.03)	0.510 (0.03)	0.454 (0.03)	0.356 (0.03)	0.308 (0.03)
University (Bachelor, Master, MBA)	0.612 (0.02)	0.666 (0.03)	0.630 (0.02)	0.578 (0.02)	0.548 (0.02)	0.535 (0.02)
University (PhD)	0.666 (0.03)	0.676 (0.04)	0.689 (0.03)	0.615 (0.03)	0.644 (0.03)	0.614 (0.04)
Experience	0.049 (0.00)	0.055 (0.00)	0.053 (0.00)	0.044 (0.00)	0.038 (0.00)	0.036 (0.00)
Experience squared $\times$ 100	-0.096 (0.00)	-0.111 (0.00)	-0.107 (0.00)	-0.088 (0.00)	-0.074 (0.00)	-0.068 (0.00)
Tenure	0.008 (0.00)	0.009 (0.00)	0.007 (0.00)	0.004 (0.00)	0.003 (0.00)	0.004 (0.00)
Tenure squared $\times$ 100	0.010 (0.00)	0.005 (0.01)	0.011 (0.00)	0.020 (0.01)	0.019 (0.01)	0.019 (0.01)
Length of interruptions	-0.001 (0.01)	0.003 (0.01)	-0.001 (0.01)	-0.011 (0.01)	-0.026 (0.01)	-0.033 (0.01)
Length of interruptions $\times$ 100	-0.064 (0.09)	-0.022 (0.10)	-0.054 (0.10)	-0.006 (0.11)	0.154 (0.11)	0.205 (0.12)
Married	0.058 (0.01)	0.060 (0.01)	0.056 (0.01)	0.044 (0.01)	0.044 (0.01)	0.046 (0.01)
Citizenship (reference group: others)						
Austrian	-0.034 (0.01)	-0.029 (0.02)	-0.017 (0.02)	-0.020 (0.02)	-0.004 (0.02)	0.018 (0.02)
EU15	0.083 (0.03)	0.044 (0.04)	0.077 (0.03)	0.039 (0.03)	0.069 (0.03)	0.140 (0.04)
Population density (reference group: high)						
Medium	-0.004 (0.01)	-0.001 (0.01)	0.002 (0.01)	-0.005 (0.01)	-0.008 (0.01)	-0.016 (0.01)
Low	-0.018 (0.01)	-0.021 (0.01)	-0.019 (0.01)	-0.020 (0.01)	-0.020 (0.01)	-0.019 (0.01)
Worker status (reference group: white collar)						
Blue collar worker	-0.092 (0.01)	-0.099 (0.01)	-0.089 (0.01)	-0.096 (0.01)	-0.112 (0.01)	-0.130 (0.01)
Civil servants	-0.006 (0.01)	-0.004 (0.02)	-0.010 (0.02)	0.001 (0.02)	0.014 (0.02)	-0.028 (0.02)
Other public sector employees	-0.115 (0.01)	-0.098 (0.01)	-0.098 (0.01)	-0.085 (0.01)	-0.125 (0.01)	-0.175 (0.01)

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Table 9 – continued from previous page

	OLS	10%	25%	50%	75%	90%
Executive position	0.092 (0.01)	0.037 (0.02)	0.063 (0.01)	0.108 (0.01)	0.133 (0.01)	0.132 (0.02)
Establishment size (reference group: 1-10)						
11-19	0.001 (0.01)	0.017 (0.02)	-0.003 (0.01)	0.015 (0.01)	-0.013 (0.01)	-0.030 (0.02)
20-49	0.036 (0.01)	0.033 (0.01)	0.027 (0.01)	0.041 (0.01)	0.023 (0.01)	0.004 (0.01)
50-499	0.059 (0.01)	0.046 (0.01)	0.041 (0.01)	0.063 (0.01)	0.050 (0.01)	0.037 (0.01)
500+	0.089 (0.01)	0.075 (0.02)	0.087 (0.01)	0.102 (0.01)	0.078 (0.02)	0.059 (0.02)
Firm size unknown	0.003 (0.02)	0.006 (0.02)	-0.008 (0.02)	-0.004 (0.02)	0.018 (0.02)	0.041 (0.02)
Log Firm size	0.028 (0.00)	0.030 (0.00)	0.029 (0.00)	0.025 (0.00)	0.023 (0.00)	0.021 (0.00)
Average age in the firm	0.004 (0.00)	0.004 (0.00)	0.004 (0.00)	0.005 (0.00)	0.005 (0.00)	0.006 (0.00)
Ratio female to male workers in the firm	-0.221 (0.02)	-0.196 (0.02)	-0.230 (0.02)	-0.256 (0.02)	-0.220 (0.02)	-0.225 (0.02)
Ratio female to male wages in the firm	-0.179 (0.02)	-0.163 (0.02)	-0.161 (0.02)	-0.145 (0.02)	-0.135 (0.02)	-0.125 (0.02)
Worker turnover in the firm	-0.000 (0.00)	0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)	-0.000 (0.00)
Number of observations	11043	11043	11043	11043	11043	11043

*Note:* Standard errors in parentheses. 11043 men in private and public sector employment. Full-time employees only. Specification 1 as depicted in Table 3 is used. All regressions include region, occupation and industry effects.